Fiscal sustainability in Romania

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Abstract. In this paper we have tested the Romanian fiscal sustainability during the period between In this way. In this regard, we analyzed the relationship between the public debt and primary budget deficit, and between public expenditure and public revenue in Romania. We aim to find out what if there are any type of relationships as Afonso and Jalles (2012) defined sustainability: unidirectional causality, bidirectional causality (perfect fiscal synchronization) and no causality. Hence, our approach was focused on testing fiscal sustainability in Romania and we have found that there was a sort of sustainability. In this period the policy makers took some measures in accordance with the definition of what fiscal sustainability should mean.

Keywords: sustainable fiscal policy, primary budget balance, public debt, public expenditures, public revenues.

JEL Classification: C22, G11.
Introduction

The economists (Blanchard et al., 1990) have argued that fiscal policy is sustainable when the public debt does not explode and the governments are not forced to raise taxes, reduce spending and increase fiscal deficit or public debt. According to them, the present value of future primary surpluses must be equal to the current level of public debt. They considered that the government should borrow money to finance the primary deficit (the difference between primary expenditure and government revenue), the payment of previous year interests and the public debt of the previous year.

In order to ensure the long-term budgetary resources as a premise of fiscal sustainability, the economic and social context that will lead to increased public spending will be taken into account.

The long-term sustainability of the public finances is also taken into account in the assessment of stability and convergence programs. In the upcoming decades, the size and age structure of Europe's population will undergo dramatic demographic changes. The aging of the population will generate great economic, budgetary and social challenges. In response, the Commission published a special report in which it concluded that if EU governments continue to implement their current policies, the public debt will increase sharply in the upcoming decades. Both fiscal consolidation, on the one hand, and general economic reforms, on the other, are needed to meet the challenge of aging. Bringing, in this way, a sustainable adjustment of the budgets, it will go a long way in improving fiscal sustainability. Bringing, in this way, a sustainable budgetary adjustment will contribute a lot to a long way in improving fiscal sustainability.

The sustainability of public finances has generated many debates that have led to the identification of different economic models developed by economists such as Gupta, Keen, Clements, Fletcher, de Mello, and Mani (2002) that involve the convergence between economic and social growth and the natural environment protection policies. Starting from this idea, the authors also highlighted the impact of fiscal policy on economic growth and investment in human capital. Corsetti and Roubini (Corsetti, 1991) has shown that policy changes or changes of relevant macroeconomic variables, such as economic growth, inflation, interest rates, are needed in the future, when budgetary constraint is not supported by empirical testing. (Moraga and Vidal, 2004) also highlighted the impact of fiscal sustainability on economic growth, as well as the need to respect the intertemporal budgetary constraint in order to ensure long-term sustainability.

Institutions such as the International Monetary Fund (IMF) and the Organization for Economic Co-operation and Development (OECD) have also focused on sustainability because, especially in the 1980s, many industrialized countries recorded significant public debt, which led to numerous episodes of fiscal adjustments in order to limit budget deficits. In addition, the member countries of the Economic and Monetary Union, as well as the new members of the European Union, are facing legal problems of tax constraints (Stoian, 2007).
Fiscal sustainability is an important element of economic sustainability, with the following components:

a) Stability – “The debt ratio target – as a percentage of GDP – must be such that, once reached as the end point of a convergence process, it can always be maintained as the average debt ratio over all subsequent economic cycles (stability)” (Hiebert, 2000).

b) Security – “must be maintained as the average debt ratio recorded in all economic cycles that have taken place after the completion of the convergence, allowing the economy to be safeguarded from any deficit situation that could be considered excessive in accordance with the rules of the Treaty” (Hiebert, 2000) which it allows a country to build its global deficit so as not to violate the 3% of GDP limit.

The conditions for fiscal sustainability, such as debt and primary surplus rates, tend to converge to their long-term benchmarks regardless of initial conditions. The issue of correlating budget revenues with the need for financial governance and obligations for a public data service that leads to the association of a sustainable fiscal policy with the solvency of government funding or clarity of liquidity. The problem of correlating the budget revenue with the need for financial governance and its obligation to serve public debt leads to the association of a sustainable fiscal policy with the solvency of government funding or clarity of liquidity. In this sense, the sustainable level of public debt is that amount of debt that could be repurchased by public authorities without adjusting revenues and expenditures in the next period. Therefore, the fiscal sustainability is strongly influenced by financing costs and the ability or willingness to satisfy the debt service.

The solvency conditions for assessing a long-term sustainable fiscal policy, formulated by Artis and Marcellino in 2002 (Artis, 2002), is that public debt coverage provides prospects for sufficient future budget surpluses. This budgetary constraint expresses an accounting identity which, according to Mendoza (2003), involves measures to meet intertemporal budgetary constraints affecting government debt. Therefore, the definition of fiscal sustainability based on intertemporal budgetary constraints allows the increase of the short- or medium-term liquidity level or long-term insolvency based on costly fiscal adjustments, according to (Roubini, 2001).

The fiscal instruments used will take into account the dynamics of economic growth, current account balance, balance of payments, budget deficits, interest rate and foreign direct investment for the forecast of available resources for the interest payments on public debt. The resources level will be compared with the value of the public debt interest and thus one could identify the type of problems that the government has to solve regarding solvency or liquidity. Consequently, the government must reduce its level of indebtedness and implement measures to restore financial discipline. However, a simple analysis of the country's indebtedness is not enough to determine the degree of sustainability of fiscal policy. Increasing the degree of indebtedness does not necessarily imply the registration of an economic growth rate higher than the real interest rate on government loans. Therefore, the solvency of the state is a necessary condition, but not sufficient for the fiscal policy (Horne, 1991) being necessary to analyze the economic environment and the origin of the financing sources for the economic environment as a whole.
(Ballabriga et al., 2005) emphasize the importance of long-term solvency, which acts as a constraint on fiscal policy, as the current public debt must be covered by the present value of future primary surpluses.

**Literature review**

The economic literature distinguishes between different definitions of financial contagion. However, the most widely used definition is that of Eichengreen et al. (1996) who consider that the contagion effect is a significant increase in the probability of crisis in one economy, conditioned by the occurrence of a crisis in another. Forbes and Rigobon (2001) have developed their own definition that states that the term contagion represents a significant increase of the links between markets, after a shock in one country or in a number of countries. The significant increase in links between financial markets involves creating or enhancing new transmission channels during the crisis, regardless of the fundamental principles and responding to a crisis in a particular country. In general, contagion refers to the spread of financial market disruptions at regional, or even global, levels.

There are several theoretical and practical studies that have focused on the analysis of the contagion phenomenon. Bekaert et al. (2014) analyzes the transmission of the financial crisis from 2007 to 2009, using a factor model to predict the return in crisis times, defining unexplained increases in factor importance and residual correlations, as an indication of contagion.

Contagion research during Sovereign Debt Crisis in Europe, based on correlation analyzes, indicates mixed results, but many of them are in the same direction. Some papers, such as Claeys and Vasicek (2014), found a significant increase in correlation coefficients between financial markets during the sovereign debt crisis in the Europe. Moreover, Horta (2013) analyzes the contagion effect of the Greek equity market on the European equity markets of the NYSE group in the context of the 2010 sovereign debt crisis, performing three contagion tests based on copula functions.

Cho and Parhizgari (2008) define and measure the contagion phenomenon by analyzing the East Asia financial crisis of 1997 on the equity markets of eight countries using DCC-GARCH.

Cho and Parhizgari (2008) define and measure the contagion phenomenon by analyzing the East Asia financial crisis of 1997 on the capital markets of eight countries using DCC-GARCH. Considering Thailand and Hong Kong as alternative sources of contagion, a total of fourteen target source pairs are analyzed using DCC-GARCH methods and median and median difference tests. They define contagion as a statistical break in the calculated conditional correlations, measured by the changes of the mean and median. It is also worth mentioning the paper of Forbes and Rigobon (2002), in which they analyze the collapse of the equity market in 1987, in addition to the 1997 crisis, with a correlation measure adjusted according to heteroskedasticity.
In addition, the paper published by Chiang et al. (2010) identified the contagion effect during the Asian crisis in 1997 by estimating a dynamic conditional correlation model (DCC). Also Syllignakis and Kouretas (2009) provides further analysis of the problem of contagion by examining correlations between seven equity markets from the CEE that have recently became EU members. This analysis was performed using the multivariate DCC GARCH model.

After discussing GARCH models that estimate contagion based on conditional correlations, we should also mention the approach developed by Diebold and Yilmaz (2009), by which they built an index that captures the volatility movement and its transmission from one market to another. Thus, this volatility spillover index is a good measure for the contagion effect, and it is also a way to quantify the impact of events on the equity market in one country and the transmission through different channels to other countries’ equity markets. Moreover, this approach used by Diebold and Yilmaz (2009) offers a possibility of exact quantification of the contagion effect between different markets, but also of the contagion effect at the level of a group, being able to observe this phenomenon in dynamics. This method of analysis allows us to draw some conclusion about how certain events have influenced the dynamics of equity markets in the recent years and the obtained results could be very easily interpreted economically.

The experience of the most recent economic and financial crisis has shown us that markets can be atypical in times of tension and thus the level of integration may be different from one period to another, depending on the general economic situation. Therefore, the level of market integration and stability may be dependent on the economic situation, but there is a possibility that the situation may also be the opposite. We can also talk about a dependence of the economic situation on the situation and stability of financial markets. One work that addressed this issue is that of Mendoza et al. (2009), but the specialized literature is quite little developed in this way.

**Methodology**

We will use in this analysis an approach according to (Afonso and Jalles, 2012) that will go in two directions of analysis in order to obtain a series of conclusions regarding fiscal sustainability. Thus, they proposed for analysis two possible relationships that could be studied through a regression and that could provide information on fiscal sustainability. We remind you that the analysis will be performed for the case of Romania, being a country for which the discussions on fiscal sustainability were not numerous in the academic environment, and the publications on this subject were quite limited.

(Afonso and Jalles, 2012) propose two types of relationships:

1) Dependent variable: primary budget balance (deficit) and independent variable: public debt (both variables will be expressed as a percentage of GDP)

\[ \text{primbugdef}_t = \alpha_0 + \alpha_1 \text{pubdebt}_t + \epsilon_t. \]
b) Dependent variable: public revenue and independent variable: public expenditure (both variables will be expressed as a percentage of GDP)

\[ \text{pubrev}_t = \alpha_0 + \alpha_1 \text{pubexp}_t + \varepsilon_t. \]

Starting from these two, we will go further towards the realization of some tests through which to verify a form of fiscal sustainability. These equations are constructed to illustrate the idea of a linear relationship between variables so that later several checks can be made using a series of concepts from the econometric literature.

First, according to (Afonso and Jalles, 2012) it can be checked the level of cointegration between the dependent variable and the independent variable, and in this way it is sought if there is a linear relationship between them, which would show that they will converge to a long-term equilibrium relationship and that there is a sustainability between these variables, i.e. between public revenues and expenditures or between the primary budget balance and public debt. Cointegration testing can be performed using the Johansen test implemented in Eviews.

Cointegration is a statistical property for a group of data series. Cointegration has become an increasingly important property in contemporary data series analysis, this concept being widely introduced by (Nelson and Plosser, 1982), noting that many time series show stochastic trend, i.e. they have unit root or are integrated of at least order 1 (I (1)). They also showed that processes that have a unit root do not have standard statistical properties and cannot be modeled with classical econometric methods. Thus, they defined the cointegration for two data series as follows: if two series are integrated by a certain order (I (n)) and a linear combination for them is integrated by a smaller order, then it can be stated that the series data are cointegrated.

The verification will be performed using the Johansen test for the data series used, two by two, according to the previously defined models. The null hypotheses of this test will be the following and will be interpreted based on calculated statistics and associated probabilities, according to (Johansen and Juselius, 1990):
1. There are no cointegration orders.
2. There is at least one cointegration order.
3. There are at least two cointegration orders.
4. There are at least 3 cointegration orders.

Thus, this testing will be performed, and the results will provide information about the level of cointegration of the data series, and they can be translated into a level of fiscal sustainability according to (Afonso and Jalles, 2012). The approach was also used in a similar way by (Gregory and Hansen, 1996) and later (Afonso and Jalles, 2012) improved and expanded it.

If it is not possible to talk about a cointegration relationship, we will try to apply a VAR (Vector Auto Regressive) model to analyze the relationship between the two variables and impulse response functions. Through these impulse response functions we can see how one of the variables reacts to a shock in another variable, and this can be very useful in our situation when we aim to test fiscal sustainability.
Next, we will move on to the second way in which I will analyze fiscal sustainability according to (Afonso and Jalles, 2012). This approach was developed by the two authors, starting from a series of works in the literature. Thus, they propose to test causality in the Granger sense for the variables considered, and the results can be put into 3 categories and can be interpreted, especially when we talk about the relationship between public revenues and expenditures:

a) One-way causality that can show us that the government adjusts expenditures or revenues according to the other variable. The same interpretation applies to the relationship between the primary deficit and public debt: the deficit or debt is adjusted taking into account the other variable, depending on the meaning of the relationship.

b) Bidirectional causality (perfect fiscal synchronization). The existence of this result verifies the classic hypothesis of public finances (Musgrave, 1966) according to which decision makers correlate expenditures and revenues and make decisions that take into account the impact on both. The same is true for the relationship between the primary budget balance (the budget balance unaffected by interest expenditure) and the share of public debt in GDP.

c) Without any causality – this situation shows us that there are no connections and that we cannot talk about a fiscal sustainability (situation identified by Hoover and Sheffrin, 1992) and later debated in a series of papers. This variant is consistent with the lack of cointegration and is the most unfavorable situation, this is the situation in which it is considered that the fiscal-budgetary policies promoted by the government are unsustainable and will lead, sooner or later, to major imbalances in the economy.

Causality testing will be performed using the Granger test implemented in Eviews, and the results will be interpreted from an econometric and statistical point of view, but also from an economic point of view.

**The data**

For quantitative analysis based on econometric methods we will use time series for the following variables:

- The share of public debt in GDP.
- Primary budget balance expressed as a share of GDP.
- Public revenues expressed as a share of GDP.
- Public expenditures expressed as a share of GDP.

All these data will be obtained for the case of Romania, using the databases from Ameco and Eurostat websites. The period for which data were obtained is represented by the period between 1998 and 2018, meaning there are 21 observations.

The data frequency is an annual one, for each of the chosen variables.

The first step in the analysis of the data series will be represented by Table 1 regarding the descriptive statistics. The result of the calculations regarding the statistical indicators established in the literature is presented below:
Table 1. Descriptive statistics

<table>
<thead>
<tr>
<th></th>
<th>Primary Budget Deficit</th>
<th>Public Debt</th>
<th>Public Expenditures</th>
<th>Public Revenues</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean</td>
<td>-1.12</td>
<td>26.20</td>
<td>36.19</td>
<td>32.90</td>
</tr>
<tr>
<td>Median</td>
<td>0.07</td>
<td>24.82</td>
<td>35.60</td>
<td>32.90</td>
</tr>
<tr>
<td>Std. dev</td>
<td>3.43</td>
<td>9.51</td>
<td>2.19</td>
<td>1.33</td>
</tr>
<tr>
<td>Kurtosis</td>
<td>0.93</td>
<td>-1.48</td>
<td>-1.04</td>
<td>-0.36</td>
</tr>
<tr>
<td>Skewness</td>
<td>-1.24</td>
<td>-0.06</td>
<td>0.28</td>
<td>-0.02</td>
</tr>
<tr>
<td>Min</td>
<td>-9.19</td>
<td>11.97</td>
<td>33.10</td>
<td>30.30</td>
</tr>
<tr>
<td>Max</td>
<td>2.63</td>
<td>39.22</td>
<td>40.00</td>
<td>35.50</td>
</tr>
<tr>
<td>Observation</td>
<td>21</td>
<td>21</td>
<td>21</td>
<td>21</td>
</tr>
</tbody>
</table>

Source: own computations.

Results

According to the methodology presented in the beginning of the case study, first of all, the cointegration testing of the two groups of data series will be performed: public expenditures and public revenues or public debt and the primary budget deficit. For this purpose, the Johansen test in Eviews will be applied.

Before testing the cointegration, it is necessary to make sure that the tested data series are integrated of the same order, and for this I will use the Augmented Dickey Fuller test applied in turn for each data series.

The first relationship to be tested will be that between public debt and the primary budget balance using the approach of authors (Afonso and Jalles, 2012).

We will test the stationarity primarily for the public debt series and for the primary budget balance, and the results are presented in the following table:

Table 2. Stationary in level for public debt and primary budget balance series (ADF)

<table>
<thead>
<tr>
<th></th>
<th>Public Debt</th>
<th>Primary Budget Balance</th>
</tr>
</thead>
<tbody>
<tr>
<td>Intercept</td>
<td>0.480</td>
<td>0.162</td>
</tr>
<tr>
<td>Trend and intercept</td>
<td>0.070</td>
<td>0.433</td>
</tr>
<tr>
<td>None</td>
<td>0.538</td>
<td>0.040</td>
</tr>
</tbody>
</table>

Source: own computations.

It can be seen that the probability for the ADF test is higher than 10% for 2 out of 3 results, and which may suggest that the data series is not stationary in level. Next, we will go and test the stationarity using the difference of order 1, meaning that we will test if the series are integrated of order 1.

Table 3. Stationary in level for public debt and primary budget balance series (ADF)

<table>
<thead>
<tr>
<th></th>
<th>Public Debt</th>
<th>Primary Budget Balance</th>
</tr>
</thead>
<tbody>
<tr>
<td>Intercept</td>
<td>0.085</td>
<td>0.090</td>
</tr>
<tr>
<td>Trend and intercept</td>
<td>0.514</td>
<td>0.266</td>
</tr>
<tr>
<td>None</td>
<td>0.028</td>
<td>0.009</td>
</tr>
</tbody>
</table>

Source: own computations.

Following the application of the ADF test using constant, constant and trend or without any of them, it was observed that the probability associated with the test is less than 10% for two of these data series, meaning that we can say this time that the series of 1st order differences are stationary and it can be said that the two data series are integrated of order 1. Therefore, it can be proceed to the cointegration test using the Johansen test according to
econometric theory. The results are presented below in the form of a summary table in Eviews for all possibilities of testing the level of cointegration.

<table>
<thead>
<tr>
<th>Selected (0.05 level*) Number of Cointegrating Relations by Model</th>
<th>Data Trend: None</th>
<th>Data Trend: Linear</th>
<th>Data Trend: Quadratic</th>
</tr>
</thead>
<tbody>
<tr>
<td>Test Type</td>
<td>No Intercept</td>
<td>Intercept</td>
<td>Intercept</td>
</tr>
<tr>
<td>No Trend</td>
<td>Trace</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>Trend</td>
<td>Max-Eig</td>
<td>0</td>
<td>0</td>
</tr>
</tbody>
</table>


Source: own computations.

It can be seen that there is a 1st order cointegration relationship, when a linear combination is made between the two variables, if constant and trend is used in the test. Also, the results of this test indicate that the data series are cointegrated and that a linear combination of the two converges to a medium to long-term equilibrium level.

Therefore, it can be stated that from the point of view of the Johansen cointegration test, the public policy promoted by the Romanian decision-makers regarding the budget balance and the public debt was a sustainable one in a certain form. This shows that there is a linear link between the two variables and that fiscal policy makers have taken into account developments in the two variables to manage public debt or the primary budget balance.

The second relationship that will be tested is the one for public expenditures and public revenues in Romania. The cointegration-based approach will be used again according to (Afonso and Jalles, 2012). As we did in the case of the previous relationship, it will be verified first of all that the data series have the same integration order using the ADF test. We will present below a summary table for the probabilities for the ADF test for the two data series using: constant, trend and constant or without either (Table 5).

| The results of the ADF test for the data series on public revenues and expenditures in Romania |
|---|---|---|
| Intercept | Public Expenditures | 0.016 |
| Trend and intercept | Public Revenues | 0.009 |
| None | Public Expenditures | 0.090 |
| None | Public Revenues | 0.041 |
| None | Public Expenditures | 0.001 |
| None | Public Revenues | 0.083 |

Source: own computations.

The results show for both data series that they are stationary in level and it is not possible to go further to check the cointegration. As an extension of the case study, a VAR model can be used here to verify the relationships between these variables and to verify how shocks are transmitted from one variable to another.

The second direction of our analysis was to test Granger causality for the two groups of data series. The results will be interpreted according to the aspects presented in the methodology part of the study, but we consider it necessary to remember which are the Granger test hypotheses:

- Null hypothesis: One variable does not cause another variable in the Granger sense.
- Alternative hypothesis: One variable causes another variable in the Granger sense.

Given these hypotheses we can say that in order to validate a causality in the Granger sense it is necessary that the probability of the test is less than the threshold of 10% to reject the
null hypothesis of the test according to which one variable does not cause another variable in the Granger sense.

The first relationship to be tested is between the primary budget balance and the share of public debt in GDP. The result of applying this test in Eviews is presented below:

<table>
<thead>
<tr>
<th>Pairwise Granger Causality Tests</th>
</tr>
</thead>
<tbody>
<tr>
<td>Lags: 2</td>
</tr>
<tr>
<td>Null Hypothesis:</td>
</tr>
<tr>
<td>PUBLIC_DEBT does not Granger Cause PRIMARY_BUDGET_DEFICIT</td>
</tr>
<tr>
<td>PRIMARY_BUDGET_DEFICIT does not Granger Cause PUBLIC_DEBT</td>
</tr>
</tbody>
</table>

Source: own computations.

Following the application of the Granger test, a causal relationship was observed from the primary budget balance to the public debt, but also a causality from the public debt to the primary budget balance. Thus, the primary budget balance influences the public debt in the Granger sense and a signal can be seen for the sustainability of public finance policies caught by this causal relationship. Also, the fact that the relationship is bidirectional shows that public debt also causes the primary budget balance in the Granger sense.

Therefore, decision-makers must take into account in the administration of public debt the primary budget balance, which will be reflected in an increase in public debt, it is necessary to finance it to cover it. Thus, a certain responsibility and fiscal sustainability at the level of public policies in Romania is highlighted.

Next, the causal relationship in the Granger sense between public revenues and public expenditures will be tested, which are expressed as a share of GDP. The result obtained is presented below:

<table>
<thead>
<tr>
<th>Pairwise Granger Causality Tests</th>
</tr>
</thead>
<tbody>
<tr>
<td>Lags: 2</td>
</tr>
<tr>
<td>Null Hypothesis:</td>
</tr>
<tr>
<td>PUB_REVENUES does not Granger Cause PUB_EXPENDITURES</td>
</tr>
<tr>
<td>PUB_EXPENDITURES does not Granger Cause PUB_REVENUES</td>
</tr>
</tbody>
</table>

Source: own computations.

Granger causality testing shows that public revenue causes public spending in the Granger sense, and the latter does not cause public revenue. Thus, the causal relationship is one-way. However, according to the results of the literature, this type of relationship indicates a certain fiscal sustainability, so decision makers in Romania are taking into account public revenues when setting the level of spending, being to some extent responsible behavior on the part of the Government and of authorizing officers when setting the budget for the following year.

This one-way causality is an approach to adjusting public finances according to Friedman (1978), in our example adjusting expenditure by income level so that revenue control leads to limited growth in the public sector: both the number of employees, as well as the level of salary or expenses with goods and services.
Conclusion

In this article we tested two relationships: the first one is the one between public debt and primary budget balance and the second one is between public expenditure and public revenue, all of these variables were computed as percentage of GDP.

After we computed the ADF test, we found out that the first relationship is not stationary, but the first differences are stationary and that the second relationship is stationarity.

Because of these ADF test results, we used Johansen test and Granger causality for first relationship and we found out that budget balance and public debt are in close contact one with the other and their evolution are correlated, a Granger causality in both way.

For the second relationship we used VAR (Vector Auto-Regressive). In this way, we tested the Granger causality and we found out that only the revenues affects the expenses, not on the other way around.

References


